

MacMahon's theorem for a set of permutations with given descent indices and right-maximal records

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Abstract

We show that the major codes and inversion codes are equidistributed over a set of permutations with prescribed descent indices and right-maximal records.

1 Introduction

Let $[n] = \{1, 2, \dots, n\}$, and let S_n be the set of permutations on $[n]$. We will use the single-line notation for a permutation: we write $\sigma = \sigma(1)\sigma(2)\cdots\sigma(n)$ rather than

$$\sigma = \begin{pmatrix} 1 & 2 & \cdots & n \\ \sigma(1) & \sigma(2) & \cdots & \sigma(n) \end{pmatrix}.$$

Given $\sigma \in S_n$, we say that i is a *descent index* of σ if $\sigma(i) > \sigma(i+1)$. We let $\text{desi}(\sigma)$ stand for the set of all descent indices of $\sigma \in S_n$. The sum of descent indices is called the *major index* of σ , denoted by $\text{maj}(\sigma)$. We say that (i, j) is an *inversion pair* if $i < j$ and $\sigma(i) > \sigma(j)$. The number of inversion pairs is referred to as the *inversion index* of σ , denoted by $\text{inv}(\sigma)$.

Let

$$E_n = \{\alpha = \alpha_1 \dots \alpha_n \mid 0 \leq \alpha_i \leq n-i, i = 1, \dots, n\}.$$

The members of E_n are called the *coding words*. Any bijective function $f : S_n \rightarrow E_n$ is referred to as *coding* of permutations. The *inversion code* is defined as

$$\text{invcode} : S_n \rightarrow E_n, \quad \text{invcode}(\sigma) = c_1 \dots c_n,$$

where c_i is the number of all inversion pairs $|\{j > i \mid \sigma(i) > \sigma(j)\}|$.

The *major code* is defined as

$$\text{majcode} : S_n \rightarrow E_n, \quad \text{majcode}(\sigma) = m_1 \dots m_n,$$

where

$$m_i = \text{maj}(\sigma^{(i)}) - \text{maj}(\sigma^{(i+1)}),$$

and $\sigma^{(i)}$ is the permutation that is obtained from σ by deleting all components less than i .

The inverse statistics

$$\text{Imajcode}, \text{Iinvcode} : S_n \rightarrow E_n$$

are defined by

$$\text{Imajcode}(\sigma) = \text{majcode}(\sigma^{-1}), \quad \text{Iinvcode}(\sigma) = \text{invcode}(\sigma^{-1}).$$

For a permutation $\sigma \in S_n$, we say that $i \in [n]$ is a *right-maximal index* of σ and that $\sigma(i)$ is a *right-maximal value*, if $\sigma(i) > \sigma(j)$ whenever $i < j \leq n$. We denote by $r[\max, i](\sigma)$ the set of all right-maximal indices of σ , and let $r[\max, v](\sigma)$ denote the set of all right-maximal values. Note that any right-maximal index is a descent index. In other words, for every $\sigma \in S_n$

$$r[\max, i] \setminus \{n\} \subseteq \text{desi}(\sigma).$$

Given a sequence α , we denote by $\text{sort}(\alpha)$ the same sequence α but written down in non-increasing order. For example, the permutation

$$\sigma = \begin{pmatrix} 1 & 2 & 3 & 4 & 5 \\ 2 & 5 & 3 & 4 & 1 \end{pmatrix} \in S_5,$$

or in our notations $\sigma = 25341$, has the descent indices 2, 4; the major index $6 = 2 + 4$; the inversion index 5; the right-maximal indices 5, 4, 2; the right-maximal values 1, 4, 5; and $\text{sort}(\alpha) = 14352$. If $\sigma = 7415236$ then $\text{invcode}(\sigma) = 6302000$, $\text{Iinvcode}(\sigma) = 2331110$, $\text{majcode}(\sigma) = 3030010$ and $\text{Imajcode}(\sigma) = 6012000$.

MacMahon [10], [11] (see also [8], [12], [9]) has proved that the major indices and the inversion indices of permutations are equidistributed over the set of all permutations,

$$|\{\sigma \in S_n \mid \text{inv}(\sigma) = k\}| = |\{\sigma \in S_n \mid \text{maj}(\sigma) = k\}|, \quad \forall k.$$

Foata [2] reproved this by constructing an explicit bijection $\phi : C \rightarrow C$, where C is the set of multiset permutations, such that $\text{maj } \sigma = \text{inv } \phi\sigma$ for every permutation $\sigma \in C$. In particular, this result holds true for usual permutation groups when $C = S_n$. Foata and Shützenberger [4] have established that the major and inversion indices are equidistributed over the set of permutations with prescribed descent indices: For any subset $A \subseteq [n-1]$,

$$|\{\sigma \in S_n \mid \text{desi}(\sigma) = A, \text{inv}(\sigma) = k\}| = |\{\sigma \in S_n \mid \text{desi}(\sigma) = A, \text{maj}(\sigma^{-1}) = k\}|, \quad \forall k.$$

It was shown in [4] that $\text{desi } \sigma = \text{desi } \phi(\sigma)$. Some further properties of ϕ were established in [1]. It was proved in particular that $r[\max, v]\sigma = r[\max, v]\phi(\sigma)$.

Hivert, Novelli, and Thibon [7] have generalized the result of [4] for major codes and inversion codes: For any subset $A \subseteq [n-1]$ and for any non-increasing coding word $\alpha \in E_n$,

$$\begin{aligned} & |\{\sigma \in S_n \mid \text{desi}(\sigma) = A, \text{sort}(\text{majcode}(\sigma^{(-1)})) = \alpha\}| \\ &= |\{\sigma \in S_n \mid \text{desi}(\sigma) = A, \text{sort}(\text{invcode}(\sigma)) = \alpha\}|. \end{aligned}$$

In our paper, the result of [7] is improved further: For any subsets A, B such that $B \setminus \{n\} \subseteq A \subseteq [n - 1]$ and for any non-increasing coding word $\alpha \in E_n$,

$$\begin{aligned} & |\{\sigma \in S_n \mid \text{desi}(\sigma) = A, \text{r}[\max, \text{i}](\sigma) = B, \text{sort}(\text{majcode}(\sigma^{(-1)})) = \alpha\}| \\ &= |\{\sigma \in S_n \mid \text{desi}(\sigma) = A, \text{r}[\max, \text{i}](\sigma) = B, \text{sort}(\text{invcode}(\sigma)) = \alpha\}|. \end{aligned}$$

Moreover, the bi-statistics $(\text{r}[\max, \text{v}], \text{majcode})$ and $(\text{r}[\max, \text{v}], \text{invcode})$ are equidistributed in a strong form (it is not necessary to sort out the majcodes and inversion codes): For any $\alpha \in E_n$,

$$\begin{aligned} & |\{\sigma \in S_n \mid \text{r}[\max, \text{v}](\sigma) = A, \text{majcode}(\sigma) = \alpha\}| \\ &= |\{\sigma \in S_n \mid \text{r}[\max, \text{v}](\sigma) = A, \text{invcode}(\sigma^{-1}) = \alpha\}|. \end{aligned}$$

Let us formulate the results of our paper in terms of generating functions.

Theorem 1.1 *The triple statistics $(\text{desi}, \text{r}[\max, \text{i}], \text{Imajcode})$ and $(\text{desi}, \text{r}[\max, \text{i}], \text{invcode})$ are equidistributed,*

$$\sum_{\sigma \in S_n} x_{\text{desi}(\sigma)} y_{\text{r}[\max, \text{i}](\sigma)} z_{\text{Imajcode}(\sigma)} = \sum_{\sigma \in S_n} x_{\text{desi}(\sigma)} y_{\text{r}[\max, \text{i}](\sigma)} z_{\text{invcode}(\sigma)}.$$

Theorem 1.2 *The bi-statistics $(\text{r}[\max, \text{v}], \text{majcode})$ and $(\text{r}[\max, \text{v}], \text{invcode})$ are non-commutative equidistributed,*

$$\sum_{\sigma \in S_n} \mathbf{x}_{\text{r}[\max, \text{v}](\sigma)} \mathbf{y}_{\text{majcode}(\sigma)} = \sum_{\sigma \in S_n} \mathbf{x}_{\text{r}[\max, \text{v}](\sigma)} \mathbf{y}_{\text{invcode}(\sigma)}.$$

Moreover, as commutative polynomials,

$$\sum_{\sigma \in S_n} x_{\text{r}[\max, \text{v}](\sigma)} y_{\text{majcode}(\sigma)} = x_n y_0 \prod_{j=1}^{n-1} (y_0 + y_1 + \cdots + y_{j-1} + x_{n-j} y_j). \quad (1)$$

Since

$$\text{r}[\max, \text{i}](\sigma^{-1}) = \text{rev}(\text{r}[\max, \text{v}](\sigma)),$$

these results can be reformulated as follows:

$$\begin{aligned} \sum_{\sigma \in S_n} x_{\text{desi}(\sigma^{-1})} y_{\text{r}[\max, \text{v}](\sigma)} z_{\text{majcode}(\sigma)} &= \sum_{\sigma \in S_n} x_{\text{desi}(\sigma^{-1})} y_{\text{r}[\max, \text{v}](\sigma)} y_{\text{invcode}(\sigma)} \cdot \\ \sum_{\sigma \in S_n} \mathbf{x}_{\text{r}[\max, \text{i}](\sigma)} \mathbf{y}_{\text{majcode}(\sigma^{-1})} &= \sum_{\sigma \in S_n} \mathbf{x}_{\text{r}[\max, \text{i}](\sigma)} \mathbf{y}_{\text{invcode}(\sigma)}. \end{aligned}$$

In fact, [7] contains one more result. They introduce one more code, the so called *saillane code*, denoted by *scode*, and proved that the bi-statistics $(\text{Idesi}, \text{majcode})$ and $(\text{Idesi}, \text{scode})$ are equidistributed as well. An extension of Hivert's result in other directions is given in [6].

There exist other kinds of permutation records. These depend on three parameters: direction (right-to-left or left-to-right), extremum (maximum or minimum) and place (index or value). Write down a permutation record briefly as $f[g, h]$, where $f = r, l$; $g = \max, \min$; and $h = i, v$. Here “r,l” corresponds to “right-to-left, left-to-right”; “max,min” to “maximum, minimum”; and “i,v” to “index, value”.

Example. If $\sigma = 516423$, then

$$l[\min, v](\sigma) = 51, l[\min, i](\sigma) = 12, l[\max, v](\sigma) = 56, l[\max, i](\sigma) = 13,$$

$$r[\min, v](\sigma) = 321, r[\min, i](\sigma) = 652, r[\max, v](\sigma) = 346, r[\max, i](\sigma) = 643.$$

The natural question appears of whether other kinds of records save equidistribution of major codes and inversion codes. We show that Theorem 1.2 cannot be improved. Changing the major (inversion) code to the saillance code is not possible. Changing the right-maximal records to other kinds of records is not possible either.

Theorem 1.3 *Let f be one of the following eight kinds of permutation records on S_n ,*

$$r[\min, i], r[\min, v], r[\max, i], r[\max, v], l[\min, i], l[\min, v], l[\max, i], l[\max, v].$$

Then the permutation bi-statistics $(f, \text{majcode})$ and $(f, \text{invcode})$ are equidistributed if and only if $f = r[\max, v]$. The bi-statistics $(f, \text{majcode}), (f, \text{scode})$ are not equidistributed.

More exactly, we establish that, if $f = r[\min, v]$ is the right-minimal values record, then

$$\sum_{\sigma \in S_n} x_{f(\sigma)} y_{\text{majcode}(\sigma)} = \sum_{\sigma \in S_n} x_{f(\sigma)} y_{\text{invcode}(\sigma^{(-1)})}$$

for $n = 2, 3, 4$ but not for $n = 5$. For the other six kinds of records, $f = r[\max, i], r[\min, i], l[\max, v], l[\max, i], l[\min, v], l[\min, i]$ and for the bi-statistics $(f, \text{majcode}), (f, \text{scode})$, counter-examples appear at $n = 3$.

2 Main Lemmas

For a coding word $\alpha = \alpha_1 \dots \alpha_n \in E_n$, we say that i is a *right-maximal index* and α_i is a *right-maximal value* of α , if $\alpha_i = n - i$.

Example. $\alpha = 140200 \Rightarrow r[\max, i](\alpha) = 642, r[\max, v](\alpha) = 024$.

Lemma 2.1 $r[\max, i](\text{invcode}(\sigma)) = r[\max, i](\sigma)$.

Proof. Let $c = \text{invcode}(\sigma) = c_1 \dots c_n$. Since $c_i \leq n - i$, c_i reaches a maximum if and only if $c_i = n - i$. Clearly, the condition $c_i = n - i$ is equivalent to the condition $\sigma(i) > \sigma(j)$ for every $j = i + 1, \dots, n$. This means that i is a right-maximal index of the coding word $c \in E_n$ if and only if i is a right-maximal index of the permutation $\sigma \in S_n$. In other words,

$$c_i = n - i \Leftrightarrow i \text{ is a right-maximal index of } \sigma .$$

•

Lemma 2.2 $r[\max, i](\text{majcode}(\sigma)) = \text{rev}(r[\max, v](\sigma))$.

Proof. Let $m = \text{majcode}(\sigma)$ and

$$r[\max, v](\sigma) = r_1 \dots r_k.$$

Recall that

$$1 \leq r_1 < r_2 < \dots < r_k = n$$

and r_i is greater than any element of σ on the right of r_i .

Let $\text{last}(\sigma^{(i)})$ be the last element of $\sigma^{(i)}$. We will look for the last elements of the sequence $\sigma^{(1)}, \dots, \sigma^{(n)}$. Let

$$\tau = \tau_1 \dots \tau_n, \quad \tau_i = \text{last}(\sigma^{(i)}).$$

Note that

$$\tau = \overbrace{\underbrace{r_1 \dots r_1}_{r_1 \text{ times}} \underbrace{r_2 \dots r_2}_{r_2 \text{ times}} \dots \underbrace{r_k \dots r_k}_{r_k \text{ times}}}$$

Therefore, $n - i$ is a descent index of $\sigma^{(i)}$ if i is a descent value of the permutation σ . In other words,

$$\text{desi}(\sigma^{(i)}) = \text{desi}(\sigma^{(i+1)}) \cup \{n - i\}$$

if and only if $i \in \text{desv}(\sigma)$. So,

$$m_i = n - i \Leftrightarrow i \text{ is a right-maximal value of } \sigma.$$

•

Example. Let $\sigma = 293785614$. Then

$$\text{invcode}(\sigma) = 171442200, \quad \text{majcode}(\sigma) = 032503010,$$

$$r[\max, i](\sigma) = 9752, r[\max, v](\sigma) = 4689.$$

We see that

$$r[\max, i](\text{majcode}(\sigma)) = 9864 = \text{rev}(r[\max, v](\sigma)),$$

$$r[\max, i](\text{invcode}(\sigma)) = 9752 = r[\max, i](\sigma).$$

Example. Let $\sigma = 86742153$. Then

$$\sigma = 86\color{red}{7}421\color{red}53 \Rightarrow r[\max, v](\sigma) = 3578$$

and

$$i \quad \sigma^{(i)} \quad \tau_i$$

1	86742153	3
2	8674253	3
3	867453	3
4	86745	5
5	8675	5
6	867	7
7	87	7
8	8	8

Therefore, the sequence of last elements is

$$\tau = 33355778.$$

Further,

$$i \quad \sigma^{(i)} \quad maj(\sigma^{(i)})$$

1	86742153	$1 + 3 + 4 + 5 + 7 = 20$
2	8674253	$1 + 3 + 4 + 6 = 14$
3	867453	$1 + 3 + 5 = 9$
4	86745	$1 + 3 = 4$
5	8675	$1 + 3 = 4$
6	867	1
7	87	1
8	8	0

Thus,

$$m_1 = 6, m_2 = 5, m_3 = 5, m_4 = 0, m_5 = 3, m_6 = 0, m_7 = 1, m_8 = 0.$$

We see that $m = 65503010$ and

$$r[\max, i](\text{majcode}(\sigma)) = 8753 = \text{rev}(r[\max, v](\sigma)).$$

Lemma 2.3 $\text{rev}(r[\max, i](\sigma)) = r[\max, v](\sigma^{-1})$.

Proof. Let $r[\max, i](\sigma) = i_1 \dots i_k$. Then

$$\sigma(i_k) = n > \sigma(i_{k-1}) > \dots > \sigma(i_1), \quad i_1 = n > i_2 > \dots > i_k.$$

Moreover, $\sigma(i_s) > \sigma(j)$ for any $i_s < j \leq n$, $s = 1, \dots, k$. Therefore,

$$r[\max, v](\sigma^{-1}) = i_k i_{k-1} \dots i_1.$$

•

In view of Lemma 2.3, Lemmas 2.1 and 2.2 can be rewritten as

$$r[\max, i](\text{majcode } \sigma) = r[\max, i](\text{Invcode}(\sigma)) = \text{rev}(r[\max, v](\sigma)) \quad (2)$$

3 Proof of Theorem 1.1

Let $A = \{a, b, c, \dots\}$ be an alphabet, A^* the set of (non-commutative) words on A , and ϵ the empty word. The *shuffle product* $w_1 \sqcup w_2$ of two words w_1 and w_2 is defined recursively by $w_1 \sqcup \epsilon = w_1$, $\epsilon \sqcup w_2 = w_2$ and

$$au \sqcup bv = a(u \sqcup bv) + b(au \sqcup v), \quad a, b \in A, u, v \in A^*.$$

For example,

$$ab \sqcup cd = abcd + acbd + acdb + cabd + cadb + cdab.$$

For a word $w = w_1 \cdots w_n$ over the integers, and $k \in \mathbf{N}$, we denote by $w[k]$ the *shifted word*

$$w[k] := (w_1 + k) \cdot (w_2 + k) \cdots (w_n + k).$$

The *shifted shuffle* of two permutations $\alpha \in S_k$ and $\beta \in S_l$ is defined by

$$\alpha \cup \beta := a \sqcup (\beta[k]).$$

A *composition* of an integer n is a sequence of positive integers of total sum n . The *descent set* $\text{Des}(I)$ of a composition $I = (i_1, \dots, i_r)$ is the set of partial sums $\{i_1, i_1 + i_2, \dots, i_1 + \cdots + i_r\}$. Compositions are ordered by $I \leqslant J$ iff $\text{Des}(I) \subseteq \text{Des}(J)$. In this case we say that I is *coarser* than J .

The *descent composition* $I = C(\sigma)$ of a permutation $\sigma \in S_n$ is the composition of n whose descents are exactly the set of descent indices of σ ,

$$\text{Des}(I) = \text{desi}(\sigma).$$

If $I = (i_1, \dots, i_r)$ is a composition of n , then we let $D_{\leqslant I}$ be the sum of all permutations each having descent composition coarser than I . Then

$$D_{\leqslant I} = (id_{i_1} \cup id_{i_2} \cup \cdots \cup id_{i_r})^\vee.$$

Here $^\vee$ is the linear involution sending each permutation to its inverse and $id_s = 12 \cdots s$ is the identity permutation of size s . The sum of all permutations whose descent composition is I will be denoted by D_I .

For example, the descent composition of the permutation $\sigma = 52413$ is $I = (1, 2, 2)$ and

$$\begin{aligned} D_{\leqslant I} &= \{12345, 21345, 31245, 41235, 51234, 12435, 21435, 31425, \\ &\quad 41325, 51324, 12534, 21534, 31524, 41523, 51423, 13425, \\ &\quad 23415, 32415, 42315, 52314, 13524, 23514, 32514, 42513, \\ &\quad 52413, 14523, 24513, 34512, 43512, 53412\}, \\ D_I &= \{21435, 21534, 31425, 31524, 32415, 32514, 41325, 41523, \\ &\quad 42315, 42513, 43512, 51324, 51423, 52314, 52413, 53412\}. \end{aligned}$$

Recall that the algebra **Sym** of noncommutative symmetric functions is the free associative algebra, on the symbol set S_n , whose basis is given by $S^I = S_{i_1} \cdots S_{i_n}$ for all compositions $I = (i_1, \dots, i_r)$ [5]. When A is an ordered alphabet, $S_n(A)$ can be realized as the sum of all nondecreasing words in A^n . The commutative image of **Sym** is the algebra of symmetric functions. The S_n are mapped to the usual complete homogeneous functions h_n .

If $I = (i_1, \dots, i_r)$ is a composition of n and $Y_n = \{y_0, y_1, \dots, y_n\}$, $Z_s = \{z_0, z_1, \dots, z_s\}$, then we denote by $\tilde{h}_k(Y_n, Z_s)$ the polynomial

$$\tilde{h}_k(Y_n, Z_s) = \sum_{0 \leq i_0 \leq i_1 \leq \dots \leq i_{k-1} < s} z_{i_0} z_{i_1} \cdots z_{i_{k-1}} + y_{n-s} \sum_{0 \leq i_0 \leq i_1 \leq \dots \leq i_{k-2} \leq i_{k-1} = s} z_{i_0} z_{i_1} \cdots z_{i_{k-2}} z_s.$$

For example,

$$\tilde{h}_3(Y_7, Z_2) = z_0^3 + z_0^2 z_1 + z_0 z_1^2 + z_1^3 + y_5(z_0^2 z_2 + z_0 z_1 z_2 + z_1^2 z_2 + z_0 z_2^2 + z_1 z_2^2 + z_2^3).$$

Lemma 3.1 *Let $I = (i_1, \dots, i_n)$ be a composition of n . Then*

$$\begin{aligned} \sum_{\sigma \in id_{i_1} \cup \dots \cup id_{i_r}} y_{r[\max, v](\sigma)} z_{\text{invcode}(\sigma)} \\ = \tilde{h}_{i_1}(Y_n, Z_{i_2+\dots+i_r}) \tilde{h}_{i_2}(Y_n, Z_{i_3+\dots+i_r}) \cdots \tilde{h}_{i_{r-1}}(Y_n, Z_{i_r}) \tilde{h}_{i_r}(Y_n, Z_0). \end{aligned}$$

Proof repeats the proof of Theorem 5.1 of [7]. We use the induction on the number of parts of I . The statement is obvious for $r = 1$. Suppose that our statement is true for the composition (i_2, \dots, i_r) . Let us prove it for I .

Let σ be an element of $id_{i_2} \cup \dots \cup id_{i_r}$ and let γ be any element in $id_{i_1} \cup \sigma$. Then $Ic_{i_1+k}(\gamma) = Ic_k(\sigma)$ for all k , where $Ic_j(\alpha)$ are the components of the inversion code of a permutation α . Moreover, the sequence $Ic_k(\gamma)$ for $k \in [1, i_1]$ is nondecreasing, since $1, \dots, i_1$ are in this order in γ , and it is bounded by the number of letters of σ ; i.e., $i_2 + \dots + i_r$. Hence, the maximum of $Ic_k(\gamma)$ is $i_2 + \dots + i_r$, and, by relation (2),

$$r[\max, v] \gamma = n - i_2 - \dots - i_r.$$

The invcode is a bijection. Therefore, no two words γ may have the same code. In particular, the first i_1 values will be different if γ runs through the elements of $id_{i_1} \cup \sigma$. On the other hand, the number of elements in $id_{i_1} \cup \sigma$ is equal to the number of nondecreasing sequences in $[0, i_2 + \dots + i_r]$. Hence all sequences appear, and

$$\sum_{\gamma \in id_{i_1} \cup \sigma} y_{r[\max, v](\gamma)} z_{\text{invcode}(\gamma)} = \tilde{h}_{i_1}(Y_n, Z_{i_2+\dots+i_r}) y_{r[\max, v](\sigma)} z_{\text{invcode}(\sigma)}.$$

•

Lemma 3.2 *Let $I = (i_1, \dots, i_n)$ be a composition of n . Then*

$$\begin{aligned} \sum_{\sigma \in id_{i_1} \cup \dots \cup id_{i_r}} y_{r[\max, v](\sigma)} z_{\text{majcode}(\sigma)} \\ = \tilde{h}_{i_1}(Y_n, Z_{i_2+\dots+i_r}) \tilde{h}_{i_2}(Y_n, Z_{i_3+\dots+i_r}) \cdots \tilde{h}_{i_{r-1}}(Y_n, Z_{i_r}) \tilde{h}_{i_r}(Y_n, Z_0) \end{aligned}$$

Proof of Lemma 3.2 repeats the proof of relation (68) of [7]. It follows from four lemmas of [7], namely Lemmas 6.2, 6.3, 6.4 and 6.5. Recall that Lemma 6.5 of [7] states the following.

Let $\beta \in S_n$ and k be an integer. The set of sorted k first components of the majcodes of the elements in $id_k \cup \beta$ is the set of all sequences $(0 \leq j_1 \leq j_2 \leq \dots \leq j_k \leq n)$. In particular, we have

$$\sum_{\sigma \in id_k \cup \beta} x_{\text{majcode}(\sigma)} = h_k(X_n) x_{\text{majcode}(\beta)}.$$

We are to specify this Lemma as follows:

$$\sum_{\sigma \in id_{i_1} \cup \dots \cup id_{i_r}} y_{r[\max, v](\sigma)} z_{\text{majcode}(\sigma)} = \tilde{h}_{i_1}(Y_n, Z_{n-i_1}) \sum_{\beta \in id_{i_2} \cup \dots \cup id_{i_r}} y_{r[\max, v](\beta)} z_{\text{majcode}(\beta)} \quad (3)$$

Let us prove this specification. For any $\beta \in id_{i_2} \cup \dots \cup id_{i_r}$ the set of the sorted i_1 first components of the majcodes of the elements in $id_{i_1} \cup \beta$ is the set of all sequences $0 \leq j_1 \leq j_2 \leq \dots \leq j_{i_1} \leq i_2 + \dots + i_r$. Therefore, the maximum in the i_1 first components of the majcodes of the elements in $id_{i_1} \cup \beta$ is $i_2 + \dots + i_r = n - i_1$. By (2) this means that the right-maximal record values of the elements in $id_{i_1} \cup \beta$ appear iff the majcodes of these elements reach the maximal value $i_2 + \dots + i_r$. •

Proof of Theorem 1.1. The claim follows from Lemmas 3.1 and 3.2. •

As in [7], Theorem 1.1 (more exactly Lemma 3.1) implies the following statement.

Corollary 3.3 *The commutative generating series for the bi-statistic $(r[\max, i], \text{invcode})$ on a descent class is given by the following determinant*

$$\tilde{r}_I(Y_n, Z_I) = \begin{vmatrix} \tilde{h}_{i_1}(Y_n, Z_{n-i_1}) & \tilde{h}_{i_1+i_2}(Y_n, Z_{n-i_1-i_2}) & \cdots & \tilde{h}_{i_1+\dots+i_r}(Y_n, Z_0) \\ 1 & \tilde{h}_{i_2}(Y_n, Z_{n-i_1-i_2}) & \cdots & \tilde{h}_{i_2+\dots+i_r}(Y_n, Z_0) \\ & & 1 & \ddots & \vdots \\ & & & \ddots & \vdots \\ & & & & 1 & \tilde{h}_{i_r}(Y_n, Z_0) \end{vmatrix}$$

4 Proof of Theorem 1.2

Since major codes and inversion codes are bijective maps, we have the inverse maps

$$\text{majcode}^{-1} : E_n \rightarrow S_n, \text{Iinvcode}^{-1} : E_n \rightarrow S_n.$$

By Lemmas 2.1, 2.2 and 2.3,

$$r[\max, v](\text{majcode}^{-1}(\alpha)) = r[\max, v](\text{Iinvcode}^{-1}(\alpha)).$$

Therefore,

$$\begin{aligned}
\sum_{\sigma \in S_n} \mathbf{x}_{r[\max, v](\sigma)} \mathbf{y}_{\text{majcode}(\sigma)} &= \sum_{\alpha \in E_n} \mathbf{x}_{r[\max, v](\text{majcode}^{-1}(\alpha))} \mathbf{y}_\alpha \\
&= \sum_{\alpha \in E_n} \mathbf{x}_{r[\max, v](\text{invcode}^{-1}(\alpha))} \mathbf{y}_\alpha \\
&= \sum_{\sigma \in S_n} \mathbf{x}_{r[\max, v](\sigma)} \mathbf{y}_{\text{invcode}(\sigma)}.
\end{aligned}$$

Suppose now that the variables $x_1, \dots, x_n, y_1, \dots, y_n$ are commutative. For a coding word $c = c_1 \dots c_n \in E_n$, set

$$\bar{c}_i = \begin{cases} i & \text{if } c_i = n - i \\ 0 & \text{otherwise} \end{cases}$$

If $c = \text{invcode}(\sigma)$ for some $\sigma \in S_n$, then by Lemma 2.1 $\bar{c}_i = i$ if and only if i is a right-maximal index of σ . Therefore,

$$\begin{aligned}
\sum_{\sigma \in S_n} x_{r[\max, i](\sigma)} y_{\text{invcode}(\sigma)} &= \sum_{c \in E_n} x_{\bar{c}_1} \cdots x_{\bar{c}_n} y_{c_1} \cdots y_{c_n} \\
&= \sum_{c_1=0}^{n-1} \sum_{c_2=0}^{n-2} \cdots \sum_{c_n=0}^0 x_{\bar{c}_1} \cdots x_{\bar{c}_n} y_{c_1} \cdots y_{c_n} \\
&= \sum_{c_1=0}^{n-1} x_{\bar{c}_1} y_{c_1} \sum_{c_2=0}^{n-2} x_{\bar{c}_2} y_{c_2} \cdots \sum_{c_n=0}^0 x_{\bar{c}_n} y_{c_n} \\
&= (x_1 y_{n-1} + \sum_{c_1=0}^{n-2} x_{\bar{c}_1} y_{c_1}) (x_2 y_{n-2} + \sum_{c_2=0}^{n-3} x_{\bar{c}_2} y_{c_2}) \cdots (x_n y_0) \\
&= (x_1 y_{n-1} + \sum_{c_1=0}^{n-2} y_{c_1}) (x_2 y_{n-2} + \sum_{c_2=0}^{n-3} y_{c_2}) \cdots (x_n y_0) \\
&= x_n y_0 (y_0 + x_{n-1} y_1) \cdots (y_0 + y_1 + \cdots + y_{n-2} + x_1 y_{n-1}).
\end{aligned}$$

Similar arguments apply to majcodes. If $m = \text{majcode}(\sigma)$ for some $\sigma \in S_n$, then by Lemma 2.2 $\bar{m}_i = i$ if and only if i is a right-maximal value of σ . Therefore,

$$\begin{aligned}
\sum_{\sigma \in S_n} x_{r[\max, v](\sigma)} y_{\text{majcode}(\sigma)} &= \sum_{m \in E_n} x_{\bar{m}_1} \cdots x_{\bar{m}_n} y_{m_1} \cdots y_{m_n} \\
&= \sum_{m_1=0}^{n-1} \sum_{m_2=0}^{n-2} \cdots \sum_{m_n=0}^0 x_{\bar{m}_1} \cdots x_{\bar{m}_n} y_{m_1} \cdots y_{m_n} \\
&= \sum_{m_1=0}^{n-1} x_{\bar{m}_1} y_{m_1} \sum_{m_2=0}^{n-2} x_{\bar{m}_2} y_{m_2} \cdots \sum_{m_n=0}^0 x_{\bar{m}_n} y_{m_n}
\end{aligned}$$

$$\begin{aligned}
&= (x_1 y_{n-1} + \sum_{m_1=0}^{n-2} x_{\bar{m}_1} y_{m_1})(x_2 y_{n-2} + \sum_{m_2=0}^{n-3} x_{\bar{m}_2} y_{m_2}) \cdots (x_n y_0) \\
&= (x_1 y_{n-1} + \sum_{m_1=0}^{n-2} y_{m_1})(x_2 y_{n-2} + \sum_{m_2=0}^{n-3} y_{m_2}) \cdots (x_n y_0) \\
&= x_n y_0 (y_0 + x_{n-1} y_1) \cdots (y_0 + y_1 + \cdots + y_{n-2} + x_1 y_{n-1}).
\end{aligned}$$

So, the bi-statistics $(r[\max, i], \text{invcode})$ and $(r[\max, v], \text{majcode})$ are equidistributed and the generating functions are given by (1).

5 Proof of Theorem 1.3

Theorem 1.3 can be reformulated as follows:

$$\sum_{\sigma \in S_n} x_{r[\max, v](\sigma)} y_{\text{majcode}(\sigma)} = \sum_{\sigma \in S_n} x_{r[\max, v](\sigma)} y_{\text{invcode}(\sigma)}.$$

In this section we show that changing right-maximal records to other kinds of records is not possible here. We prove the following result:

The permutation statistics $(f, \text{majcode})$ and $(f, \text{Iinvcode})$ are not equidistributed if $f = l[a, b]$, $a = \min, \max$, $b = i, v$ or $f = r[\max, i], r[\min, i], r[\min, v]$.

Consider the test functions

$$test_1(k, f) = \sum_{\sigma \in S_k} x_{f(\sigma)} y_{\text{invcode}(\sigma^{-1})}, \quad test_2(k, f) = \sum_{\sigma \in S_k} x_{f(\sigma)} y_{\text{majcode}(\sigma)},$$

$$test(k, f) = test_1(k, f) - test_2(k, f).$$

To simplify calculations, set $x_0 = y_0 = 1$.

For $n = 3$, the following relations hold:

$$test_1(3, l[\min, v]) = x_1 + x_1 y_1 + x_1 x_2 y_1 + x_1 x_3 y_1^2 + x_1 x_2 y_2 + x_1 x_2 x_3 y_1 y_2,$$

$$test_2(3, l[\min, v]) = x_1 + x_1 x_2 y_1 + x_1 x_3 y_1 + x_1 y_1^2 + x_1 x_2 y_2 + x_1 x_2 x_3 y_1 y_2,$$

$$test(3, l[\min, v]) = x_1 y_1 (-1 + x_3) (-1 + y_1),$$

$$test_1(3, l[\min, i]) = x_1 + x_1 y_1 + x_1 x_2 y_1 + x_1 x_2 y_1^2 + x_1 x_3 y_2 + x_1 x_2 x_3 y_1 y_2,$$

$$test_2(3, l[\min, i]) = x_1 + 2x_1 x_2 y_1 + x_1 y_1^2 + x_1 x_3 y_2 + x_1 x_2 x_3 y_1 y_2,$$

$$test(3, l[\min, i]) = x_1 y_1 (-1 + x_2) (-1 + y_1),$$

$$test_1(3, l[\max, v]) = x_1 x_2 x_3 + x_1 x_3 y_1 + x_2 x_3 y_1 + x_3 y_1^2 + x_2 x_3 y_2 + x_3 y_1 y_2,$$

$$test_2(3, l[\max, v]) = x_1 x_2 x_3 + x_3 y_1 + x_2 x_3 y_1 + x_1 x_3 y_1^2 + x_2 x_3 y_2 + x_3 y_1 y_2,$$

$$test(3, l[\max, v]) = -y_1 (-1 + x_1) x_3 (-1 + y_1),$$

$$\begin{aligned}
test_1(3, l[\max, i]) &= x_1x_2x_3 + x_1x_2y_1 + x_1x_3y_1 + x_1y_1^2 + x_1x_2y_2 + x_1y_1y_2, \\
test_2(3, l[\max, v]) &= x_1x_2x_3 + x_1y_1 + x_1x_3y_1 + x_1x_2y_1^2 + x_1x_2y_2 + x_1y_1y_2, \\
test(3, l[\max, v]) &= -x_1y_1(-1 + x_2)(-1 + y_1), \\
test_1(3, r[\max, i]) &= x_3 + x_3y_1 + x_2x_3y_1 + x_1x_3y_1^2 + x_2x_3y_2 + x_1x_2x_3y_1y_2, \\
test_2(3, r[\max, i]) &= x_3 + x_3y_1 + x_1x_3y_1 + x_2x_3y_1^2 + x_2x_3y_2 + x_1x_2x_3y_1y_2, \\
test(3, r[\max, i]) &= y_1(x_1 - x_2)x_3(-1 + y_1), \\
test_1(3, r[\min, v]) &= x_1x_2x_3 + x_1x_2y_1 + x_1x_3y_1 + x_1x_2y_1^2 + x_1y_2 + x_1y_1y_2, \\
test_2(3, r[\min, v]) &= x_1x_2x_3 + x_1x_2y_1 + x_1x_3y_1 + x_1x_2y_1^2 + x_1y_2 + x_1y_1y_2, \\
test(3, r[\min, v]) &= 0, \\
test_1(3, r[min, i]) &= x_1x_2x_3 + x_1x_3y_1 + x_2x_3y_1 + x_2x_3y_1^2 + x_3y_2 + x_3y_1y_2, \\
test_2(3, r[min, i]) &= x_1x_2x_3 + 2x_2x_3y_1 + x_1x_3y_1^2 + x_3y_2 + x_3y_1y_2, \\
test(3, r[min, i]) &= -y_1(x_1 - x_2)x_3(-1 + y_1).
\end{aligned}$$

So, in the six cases $f = l[\min, i], l[\min, v], l[\max, i], l[\max, v], r[\min, i], r[\max, i]$ counter-examples appear at $n = 3$.

For $f = r[\min, v]$ and $n = 4$, no counter-examples exist,

$$\begin{aligned}
test_1(4, r[\min, v]) &= x_1x_2x_3x_4 + x_1x_2x_3y_1 + x_1x_2x_4y_1 + x_1x_3x_4y_1 + x_1x_3y_1^2 \\
&\quad + x_1x_2x_3y_1^2 + x_1x_2x_4y_1^2 + x_1x_2x_3y_1^3 + x_1x_2y_2 \\
&\quad + x_1x_4y_2 + 2x_1x_2y_1y_2 + x_1x_3y_1y_2 + x_1x_4y_1y_2 \\
&\quad + x_1x_2y_1^2y_2 + x_1x_3y_1^2y_2 + x_1x_2y_2^2 + x_1x_2y_1y_2^2 \\
&\quad + x_1y_3 + 2x_1y_1y_3 + x_1y_1^2y_3 + x_1y_2y_3 + x_1y_1y_2y_3, \\
test_1(4, r[\min, v]) &= test_2(4, r[\min, v]), \\
test(4, r[\min, v]) &= test_1(4, r[\min, v]) - test_2(4, r[\min, v]) = 0.
\end{aligned}$$

A counter-example for $f = r[\min, v]$ appears at $n = 5$. Let us prove this.

Let

$$\begin{aligned}
M &= \{\sigma \in S_5 | r[\min, v](\sigma) = 431\}, \\
M_1 &= \{\sigma^{-1} \in S_5 | r[\min, v](\sigma) = 431\}.
\end{aligned}$$

Note that

$$\begin{aligned}
M &= \{21354, 21534, 25134, 52134\}, \\
M_1 &= \{21354, 21453, 31452, 32451\}.
\end{aligned}$$

Let

$$\begin{aligned}
majcode(M) &= \{majcode(\sigma) | \sigma \in M\}, \\
invcode(M_1) &= \{invcode(\sigma) | \sigma \in M_1\}.
\end{aligned}$$

Then

$$\begin{aligned}\text{majcode}(M) &= \{21110, 21010, 01010, 20010\}, \\ \text{invcode}(M_1) &= \{10010, 10110, 20110, 21110\}.\end{aligned}$$

Hence

$$20010 \in \text{majcode}(M), \quad 20010 \notin \text{invcode}(M_1).$$

Moreover, there exists exactly one permutation $\sigma \in S_5$ such that

$$r[\min, v](\sigma) = 431, \quad \text{majcode}(\sigma) = 20010,$$

(namely, $\sigma = 52134$), but there is no permutation $\sigma \in S_5$ with the properties

$$r[\min, v](\sigma) = 431, \quad y_{\text{invcode}(\sigma^{-1})} = y_1y_2.$$

So, we have established that the sum $\sum_{\sigma \in S_5} x_{r[\min, v](\sigma)} y_{\text{majcode}(\sigma)}$ contains the member $x_1x_3x_4y_1y_2$ with coefficient 1, whereas the sum $\sum_{\sigma \in S_5} x_{r[\min, v](\sigma)} y_{\text{invcode}(\sigma^{-1})}$ does not; a contradiction.

Similarly, one can check that $\text{test}(3, f) \neq 0$ for a test function defined by

$$\text{test}(k, f) = \sum_{\sigma \in S_k} x_{f(\sigma)} y_{\text{scode}(\sigma)} - \sum_{\sigma \in S_k} x_{f(\sigma)} y_{\text{majcode}(\sigma)}.$$

Remark. We say that two triple statistics (f, g, h) and (f_1, g_1, h_1) are equidistributed, and write $(f, g, h) \sim (f_1, g_1, h_1)$, if their multi-variable generating functions are equal,

$$\sum_{\sigma \in S_n} x_{f(\sigma)} y_{g(\sigma)} z_{h(\sigma)} = \sum_{\sigma \in S_n} x_{f_1(\sigma)} y_{g_1(\sigma)} z_{h_1(\sigma)}.$$

Let $\text{Idesi}(\sigma) = \text{desi}(\sigma^{-1})$. One can show that the following triple statistics are equidistributed in a weaker form:

$$(\text{Idesi}, l[\max, v], \text{Iinvcode}) \sim (\text{Idesi}, l[\max, v], \text{majcode})$$

if $y_i = 1, i < n - 1$,

$$(\text{Idesi}, l[\min, v], \text{Iinvcode}) \sim (\text{Idesi}, l[\min, v], \text{majcode})$$

if $y_i = 1, i > 2$,

$$(\text{Idesi}, r[\min, v], \text{Iinvcode}) \sim (\text{Idesi}, r[\min, v], \text{majcode})$$

if $y_i = 1, 2 < i < n - 1$,

$$(\text{Idesi}, l[\min, v], \text{scode}) \sim (\text{Idesi}, l[\min, v], \text{majcode})$$

if $y_i = 1, i > 2$,

$$(\text{Idesi}, r[\min, v], \text{scode}) \sim (\text{Idesi}, r[\min, v], \text{majcode})$$

if $y_i = 1, i > 2$,

$$(\text{Idesi}, \text{l}[\max, \text{v}], \text{scode}) \sim (\text{Idesi}, \text{l}[\max, \text{v}], \text{majcode})$$

if $y_i = 1, i < n - 1$,

$$(\text{Idesi}, \text{r}[\max, \text{v}], \text{scode}) \sim (\text{Idesi}, \text{r}[\max, \text{v}], \text{majcode})$$

if $y_i = 1, i < n - 1$.

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